

## Reverse-modeling of biological networks

### Introduction to ODEs and state-space models

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# Outline

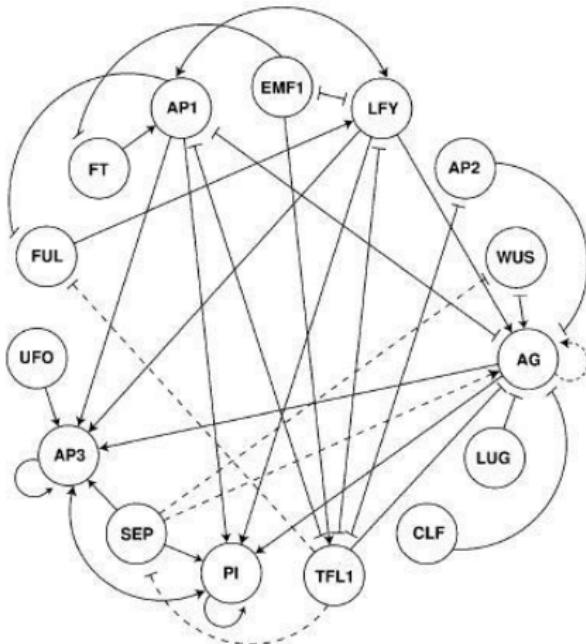
## 1 Introduction

## 2 Dynamical models

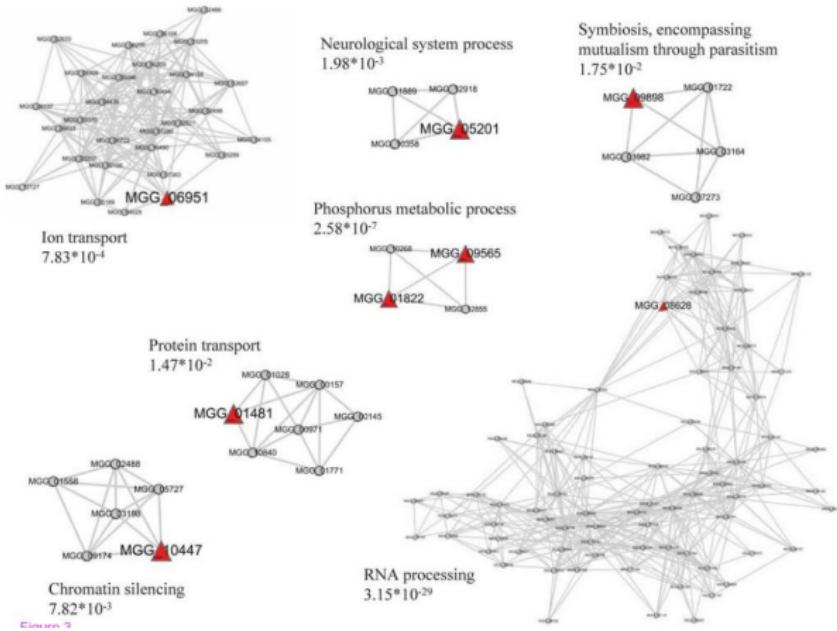
## 3 Dynamical graphical models

- Parameter estimation

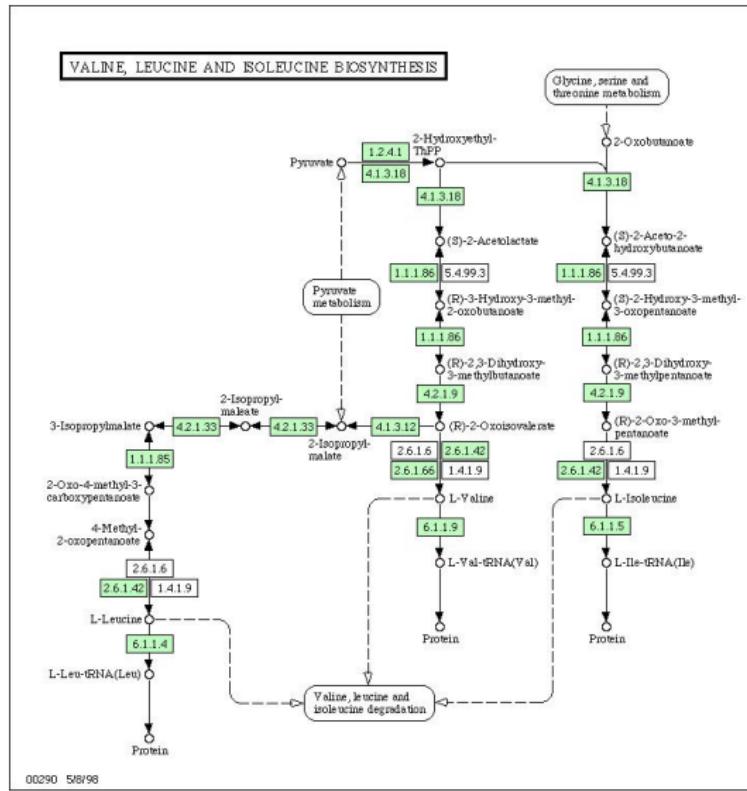
# Gene regulatory networks



# Protein-protein interaction networks (rice blast fungus)



# Metabolic networks



# Motivation

- Networks:
  - signaling pathways
  - gene regulatory networks
  - protein-protein interaction networks
  - metabolic pathways
- Goals:
  - Identify and understand complex mechanisms at work in the cell
  - Potential outputs: therapeutical targetting
- Means:
  - Use experimental data to identify networks: here comes machine learning

# How to learn biological networks from data ?

- **Predictive approaches** : predict (only) edges in an unsupervised or supervised way (scale ( $10^3$ ) nodes)
- **Modeling approaches** : model the network as a complex system and identify it (scale : 10 to a few  $10^2$  nodes)

## Motivation for modeling approaches

- A model reflects assumptions about the behaviour of the network as a dynamical system
- Once learnt, a model can be used in simulation and prediction
- Be careful, a model IS NOT the truth, it is just a tool to test hypothesis

# Biological networks as dynamical systems

## Two families of models

- ① **Differential Equations:** bifurcation analysis, control theory, theory of dynamical systems
- ② **Dynamical Graphical models:** frequentist and Bayesian approaches, incomplete data, building blocks in mixture models (switching models, modular approaches)

# System of Ordinary Differential Equations (ODE)

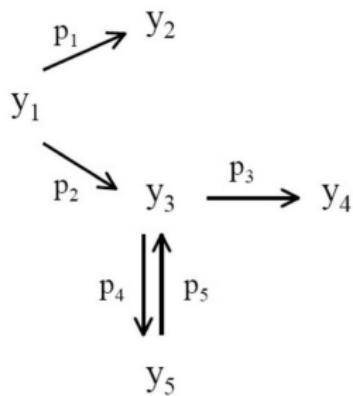
$$\frac{d\mathbf{x}}{dt} = \mathbf{f}(\mathbf{x}(t), \mathbf{u}(t); \theta)$$

- gene regulatory networks
- $\mathbf{x}(t)$  : state variables at time  $t$ 
  - $\mathbf{p}(t)$  : protein concentrations
  - $\mathbf{r}(t)$  : mRNA concentrations
- $\mathbf{f}$  : the form of  $f$  encodes the nature of interactions (and their structure)
  - linear/nonlinear models
  - Michaelis-Menten kinetics
  - Mass action kinetics
  - ...
- $\theta$ : parameter set (kinetic parameters, rate constants,...)
- $\mathbf{u}(t)$ : input variables at time  $t$

# Examples of ODEs

- Isomerization of  $\alpha$ -pinene
- JAK-STAT pathways
- Repressilator

## Isomerization of $\alpha$ -pinene network



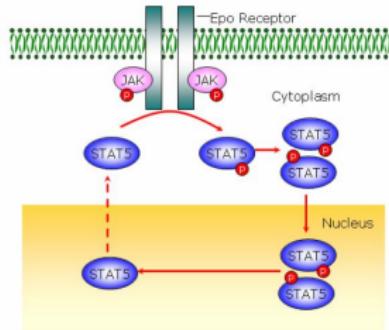
- ODEs:

$$\begin{aligned}
 \dot{x}_1 &= -(p_1 + p_2)x_1 \\
 \dot{x}_2 &= p_1 x_1 \\
 \dot{x}_3 &= p_2 x_1 - (p_3 + p_4)x_3 + p_5 x_5 \\
 \dot{x}_4 &= p_3 x_3 \\
 \dot{x}_5 &= p_4 x_3 + p_5 x_5
 \end{aligned}$$

- Isomerization of  $\alpha$ -pinene ( $y_1$ ) to dipentene ( $y_2$ ) and allo-ocimen ( $y_3$ )
- which in turn yields  $\alpha$ - and  $\beta$ -pyronene ( $y_4$ ) and a dimer ( $y_5$ )

# JAK-STAT signaling pathway

[Swameye, PNAS 2003]



- ODEs:

$$\begin{cases} \dot{x}_1(t) &= -a_1 x_1(t)u(t) + 2a_4 x_4(t)1_{\{t \geq \tau\}} \\ \dot{x}_2(t) &= a_1 x_1(t)u(t) - 2a_4 x_2^2(t) \\ \dot{x}_3(t) &= -a_3 x_3(t) + x_2^2(t) \\ \dot{x}_4(t) &= a_3 x_3(t) - a_4 x_4(t)1_{\{t \geq \tau\}} \end{cases}$$

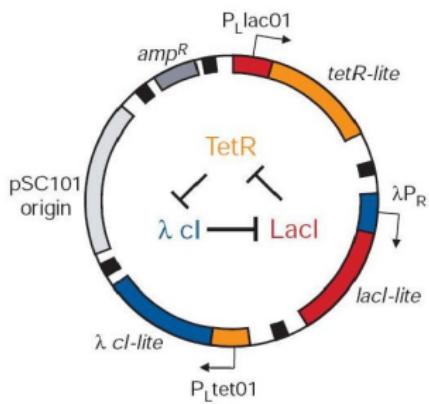
- Observed variables

$$y_1 = x_2 + 2x_3$$

$$y_2 = x_1 + x_2 + 2x_3$$

- Experimental data: 16 time points
- $\theta = (a_1, a_3, a_4)^\top$  is the parameters to be estimated

# Repressilator (Elowitz and Leibler 2000)



$$\frac{dr_1}{dt} = v_1^{\max} \frac{k_{12}^n}{k_{12}^n + p_2^n} - k_1^{\text{mRNA}} r_1$$

$$\frac{dr_2}{dt} = v_2^{\max} \frac{k_{23}^n}{k_{23}^n + p_3^n} - k_2^{\text{mRNA}} r_2$$

$$\frac{dr_3}{dt} = v_3^{\max} \frac{k_{31}^n}{k_{31}^n + p_1^n} - k_3^{\text{mRNA}} r_3$$

$$\frac{dp_1}{dt} = k_1 r_1 - k_1^{\text{protein}} p_1$$

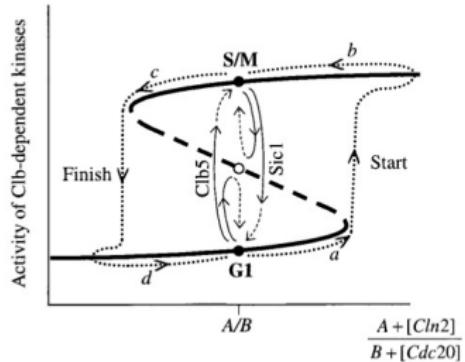
$$\frac{dp_2}{dt} = k_2 r_2 - k_2^{\text{protein}} p_2$$

$$\frac{dp_3}{dt} = k_3 r_3 - k_3^{\text{protein}} p_3$$

- mRNAs are observed, proteins are hidden
- mRNA and protein degradation rate constants are supposed to be known
- Estimate 9 parameters

# ODE: analysis of the behavior

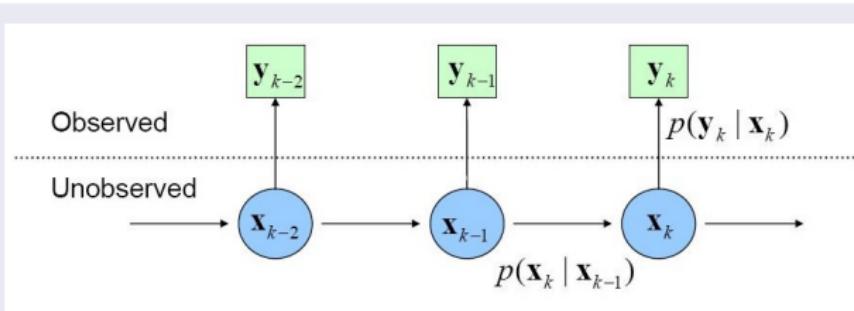
## Bistability and hysteresis (example): budding yeast



# Graphical models : from simple interactions models to complex ones

- Graphical Gaussian Model model : estimating partial correlation as a measure of conditional independency
- Bayesian networks : modeling directed interactions
- Dynamic Bayesian Networks : modeling directed interactions through time
- **State-space models** : modeling observed and hidden dynamical processes as well

## State-space models



## Equations



$$\mathbf{x}_{t+1} = \mathbf{F}_t \mathbf{x}_t + \mathbf{w}_t \quad (1)$$

$$\mathbf{y}_t = \mathbf{H}_t \mathbf{x}_t + \mathbf{v}_t \quad (2)$$

where:

- $\mathbf{F}_t$  is the transition matrix,  $\mathbf{H}_t$  is the measurement matrix
- $\mathbf{w}_t \sim N(0, \Sigma_w)$ ,  $\mathbf{v}_t \sim N(0, \Sigma_v)$
- and  $\mathbf{x}_0 \sim N(\mathbf{x}_{0|-1}, P_{0|-1})$ .

## What to do with a state-space model and a dataset ?

- Infer hidden states (given parameters)
- Estimate unknown parameters: a learning problem that requires hidden states inference

# Hidden state estimation: Kalman filtering and smoothing

- The problem of state estimation is addressed in the Bayesian framework
- Compute the filtering distribution  $p(\mathbf{x}_t | \mathbf{y}_{0:t})$
- Method : recursive estimation
- Kalman's proposal (1960) is to recursively estimate these states using the linearity of the model and the gaussianity of distributions at hand
- Various estimates can be built from the filtering distribution: MMSE, MAP (see for instance Chen's review on filtering)

# Minimum Mean Squared Error estimate

- Find  $\hat{\mathbf{x}}_t$  that minimizes

$$\int (\mathbf{x}_t - \hat{\mathbf{x}}_t)^T (\mathbf{x}_t - \hat{\mathbf{x}}_t) p(\mathbf{x}_t | y_{0:t}) d\mathbf{x}_t$$

- Solution : mean posterior  $E[x_t | y_{0:t}]$
- Also estimate error covariance matrix

# Bayesian optimal filtering

- Suppose that the parameters are known, recursively calculate the filtering distribution of the states  $p(\mathbf{x}_t | \mathbf{y}_{0:t})$

## Two steps

① Prediction:  $p(\mathbf{x}_{t+1} | \mathbf{y}_{0:t}) = \int p(\mathbf{x}_{t+1} | \mathbf{x}_t) p(\mathbf{x}_t | \mathbf{y}_{0:t}) d\mathbf{x}_t$

② Update:

$$p(\mathbf{x}_{t+1} | \mathbf{y}_{0:t+1}) = \frac{p(\mathbf{y}_{t+1} | \mathbf{x}_{t+1}) p(\mathbf{x}_{t+1} | \mathbf{y}_{0:t})}{p(\mathbf{y}_{t+1} | \mathbf{y}_{0:t})}$$

where:

$$p(\mathbf{y}_{t+1} | \mathbf{y}_{0:t}) = \int p(\mathbf{y}_{t+1} | \mathbf{x}_{t+1}) p(\mathbf{x}_{t+1} | \mathbf{y}_{0:t}) d\mathbf{x}_{t+1}$$

# Derivation of Bayesian optimal filtering for the linear case (1)

- Assuming that the posterior distribution at previous step  $p(\mathbf{x}_t|\mathbf{y}_{0:t})$  is a Gaussian density  $N(\mathbf{x}_t|\mathbf{m}_t, P_t)$
- Then  $p(\mathbf{x}_{t+1}|\mathbf{x}_t)$  is a Gaussian density of the following form:  $N(\mathbf{x}_{t+1}|F_t\mathbf{m}_t, \Sigma_w)$  using SSM equations
- How to compute  $p(\mathbf{x}_{t+1}|\mathbf{y}_{0:t})$  ?
  - $p(\mathbf{x}_{t+1}|\mathbf{y}_{0:t}) = \int p(\mathbf{x}_{t+1}|\mathbf{x}_t)p(\mathbf{x}_t|\mathbf{y}_{0:t})d\mathbf{x}_t$
  - Gaussianity of the two distributions gives the prediction step:
  - $p(\mathbf{x}_{t+1}|\mathbf{y}_{1:t}) = N(\mathbf{x}_{t+1}|F_t\mathbf{m}_t, F_tP_tF_t^T + \Sigma_w)$
  - Mean:  $\mathbf{m}_{t+1|t} = F_t\mathbf{m}_t$
  - Covariance:  $P_{t+1|t} = F_tP_tF_t^T + \Sigma_w$

# Derivation of Bayesian optimal filtering for the linear case (2)

- We need to compute:  $p(\mathbf{y}_{t+1}|\mathbf{x}_{t+1})p(\mathbf{x}_{t+1}|\mathbf{y}_{0:t})$
- $p(\mathbf{y}_{t+1}|\mathbf{x}_{t+1})p(\mathbf{x}_{t+1}|\mathbf{y}_{0:t}) = p(\mathbf{x}_{t+1}, \mathbf{y}_{t+1}|\mathbf{y}_{0:t})$
- Using Gaussianity, we can write:

$$p(\mathbf{x}_{t+1}, \mathbf{y}_{t+1}|\mathbf{y}_{1:t}) = N([\mathbf{x}_{t+1} \mathbf{y}_{t+1}]^T, \mathbf{m}'', \mathbf{P}'')$$

- $\mathbf{m}'' = [\mathbf{m}_{t+1|t}, H_{t+1} \cdot \mathbf{m}_{t+1|t}]^T$

- Block matrix:  $\mathbf{P}'' = [\mathbf{P}_{t+1|t}, P_{t+1|t} \cdot H_{t+1}; H_{t+1} \cdot P_{t+1|t}, H_{t+1} \cdot P_{t+1|t} H_{t+1}^T + \Sigma_v]$

## Reminder (gaussianity composition rule)

$$N(\mathbf{x} | \mathbf{m}, \mathbf{P}) = \frac{1}{(2\pi)^{n/2} |\mathbf{P}|^{1/2}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mathbf{m})^T \mathbf{P}^{-1} (\mathbf{x} - \mathbf{m})\right)$$

Let  $\mathbf{x}$  and  $\mathbf{y}$  have the Gaussian densities

$$p(\mathbf{x}) = N(\mathbf{x} | \mathbf{m}, \mathbf{P}), \quad p(\mathbf{y} | \mathbf{x}) = N(\mathbf{y} | \mathbf{Hx}, \mathbf{R}),$$

Then the joint and marginal distributions are

$$\begin{aligned} \begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix} &\sim N\left(\begin{pmatrix} \mathbf{m} \\ \mathbf{Hm} \end{pmatrix}, \begin{pmatrix} \mathbf{P} & \mathbf{PH}^T \\ \mathbf{HP} & \mathbf{HPH}^T + \mathbf{R} \end{pmatrix}\right) \\ \mathbf{y} &\sim N(\mathbf{Hm}, \mathbf{HPH}^T + \mathbf{R}). \end{aligned}$$

# Derivation of Bayesian optimal filtering for the linear case (3)

## Adding a new observation

- We want to compute:  $p(\mathbf{x}_{t+1} | \mathbf{y}_{0:t+1}) = N(\mathbf{x}_{t+1} | \mathbf{m}_{t+1}, P_{t+1})$
- To obtain  $p(\mathbf{x}_{t+1} | \mathbf{y}_{0:t+1})$ , we notice that:
  - $p(\mathbf{x}_{t+1}, \mathbf{y}_{t+1} | \mathbf{y}_{0:t}) = p(\mathbf{x}_{t+1} | \mathbf{y}_{0:t+1})$
- As we have already computed  $p(\mathbf{x}_{t+1}, \mathbf{y}_{t+1} | \mathbf{y}_{0:t})$ ,
- We compute the conditional distribution of  $\mathbf{x}_{t+1}$  given  $\mathbf{y}_{t+1}$  using rule 2 (gaussianity again)
- mean update:

$$\mathbf{m}_{t+1} = \mathbf{m}_{t+1|t} + K_{t+1} (\mathbf{y}_{t+1} - H_{t+1} \cdot \mathbf{m}_{t+1|t})$$

- covariance update:

$$\mathbf{P}_{t+1} = \mathbf{P}_{t+1|t} - K_{t+1} S_{t+1} K_{t+1}^T$$

- with the following notation:
- 

$$S_{t+1} = H_{t+1} \cdot P_{t+1|t} H_{t+1}^T + \Sigma_v$$

## Reminder (gaussianity composition rule)

If the random variables  $\mathbf{x}$  and  $\mathbf{y}$  have the joint Gaussian probability density

$$\begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix} \sim N \left( \begin{pmatrix} \mathbf{a} \\ \mathbf{b} \end{pmatrix}, \begin{pmatrix} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^T & \mathbf{B} \end{pmatrix} \right),$$

Then the marginal and conditional densities of  $\mathbf{x}$  and  $\mathbf{y}$  are given as follows:

$$\mathbf{x} \sim N(\mathbf{a}, \mathbf{A})$$

$$\mathbf{y} \sim N(\mathbf{b}, \mathbf{B})$$

$$\mathbf{x} | \mathbf{y} \sim N(\mathbf{a} + \mathbf{C} \mathbf{B}^{-1} (\mathbf{y} - \mathbf{b}), \mathbf{A} - \mathbf{C} \mathbf{B}^{-1} \mathbf{C}^T)$$

$$\mathbf{y} | \mathbf{x} \sim N(\mathbf{b} + \mathbf{C}^T \mathbf{A}^{-1} (\mathbf{x} - \mathbf{a}), \mathbf{B} - \mathbf{C}^T \mathbf{A}^{-1} \mathbf{C}).$$

# Recursive equations of Kalman filter

- Initialization

- Prediction step

- $\bullet \quad \mathbf{m}_{t+1|t} = F_t \mathbf{m}_t$
- $\bullet \quad P_{t+1|t} = F_t P_t F_t^T + \Sigma_w$

- Update step

- $\bullet \quad \mathbf{z}_{t+1} = \mathbf{y}_{t+1} - H_{t+1} \mathbf{m}_{t+1|t}$
- $\bullet \quad S_{t+1} = H_{t+1} \cdot P_{t+1|t} H_{t+1}^T + \Sigma_v$
- $\bullet \quad K_{t+1} = P_{t+1|t} \cdot H_{t+1}^T \cdot S_{t+1}^{-1}$
- $\bullet \quad \mathbf{m}_t = \mathbf{m}_{t+1|t} + K_{t+1} \mathbf{z}_{t+1}$
- $\bullet \quad P_{t+1} = P_{t+1|t} - K_{t+1} S_{t+1} K_{t+1}^T.$

## Estimation of hidden state (linear case)

- Kalman Filter computes the Gaussian posterior probability by computing its mean and covariance estimates
- We thus get the MMSE estimate

# Improvement by smoothing

- Filtering provides an estimate of the posterior probability  $p(\mathbf{x}_t|y_0, \dots, y_t)$
- Smoothing provides an estimate of the posterior probability  $p(\mathbf{x}_t|y_0, \dots, y_T)$
- Recursive equations of smoothing can be derived similarly to filtering
- Smoothing is applied backward after filtering

# Parameter learning

Several approaches among which:

- ① Maximum likelihood principle with EM algorithm (also can be used for MAP)
- ② Augmented approach : parameters as stationnary state (for instance Sitz et al. 2002, Quach et al. 2007)

# Dealing with nonlinear state-space models

## Nonlinear model



$$\begin{aligned}\mathbf{x}(t+1) &= \mathbf{F}(t, \mathbf{x}(t); \theta) \\ \mathbf{y}(t) &= \mathbf{H}(t, \mathbf{x}(t); \theta) + \epsilon(t)\end{aligned}$$

- When  $\mathbf{F}, \mathbf{H}$  are nonlinear, the integrals are usually intractable.  
Approximate solutions are needed!
- Issue : estimate sufficient statistics of  $p(\mathbf{x}_t | \mathbf{y}_{0:t})$ 
  - Extended Kalman Filter
  - Unscented Kalman Filter
  - (sequential) Monte-Carlo strategies

# Unscented Kalman Filter

- S. J. Julier and J. K. Uhlmann, A New Extension of the Kalman Filter to Nonlinear Systems, 1997.
- Eric A. Wan and Rudolph van der Merwe, The Unscented Kalman Filter for Nonlinear Estimation, 2000.
- Extension of UKF to a broader class of nonlinear estimation problems

# Unscented Kalman Filter

## Issues

- Estimate  $\mathbf{m}_t, P_t, \mathbf{m}_{t+1|t}, P_{t+1|t}$  while transformations are nonlinear
- UKF is based on the Unscented Transform (UT)
- UT is a method for calculating the statistics of a random variable which undergoes a nonlinear transformation

## Unscented transform

- Consider  $f$ , a nonlinear transform of a random variable  $X : Y = f(X)$
- Given  $\bar{X} = E[X]$  and  $P_X = E[(X - \bar{X})(X - \bar{X})]$ ,
- Find:  $\bar{Y} = E[Y]$  and  $P_Y = E[(Y - \bar{Y})(Y - \bar{Y})]$

# Unscented Transform

## UT algorithm (Julier, 1996)

- ➊ here  $d$  is the dimension of the variable to approximate
- ➋ Compute  $2d + 1$  sigma points:  $\{\xi_i\} = \bar{X} + (\alpha\sqrt{P_X})_i, i = 1\dots d$  and  $\{\xi_i\} = \bar{X} - (\alpha\sqrt{P_X})_i, i = d + 1\dots 2d + 1$
- ➌ Transform each sigma point:  $\psi_i = f(\xi_i)$
- ➍ Compute a weighted sample mean and covariance:
- ➎  $\bar{Y} = \sum_{i=0}^{2d} \beta_i \psi_i$  and  $P_Y = \sum_{i=0}^{2d} \beta_i (\psi_i - \bar{Y})(\psi_i - \bar{Y})^\top$

# Illustration

Figure: Unscented Transform (Wan's paper)